



Working Paper 2016-04

Taxation of Temporary Jobs: Good Intentions With Bad Outcomes ?

Pierre Cahuc
Olivier Charlot
Franck Malherbet
Hélène Benghalem
Emeline Limon

Taxation of Temporary Jobs: Good Intentions With Bad Outcomes ?

Pierre Cahuc^{1 2}

CREST, Ecole Polytechnique, IZA, CEPR

Olivier Charlot³

Université Cergy-Pontoise, THEMA

Franck Malherbet⁴

Université de Rouen, Ecole Polytechnique, CREST, IZA

Helène Benghalem⁵

CREST, Ecole Polytechnique

Emeline Limon⁶

Université Cergy-Pontoise, THEMA

May 2016

¹We thank the *Chaire Sécurisation des Parcours Professionnels* for its support and *Unédic* for providing data. We also thank Normann Rion for excellent research assistance.

²Contact: pierre.cahuc@polytechnique.edu.

³Contact: ocharlot@u-cergy.fr

⁴Contact: franck.malherbet@univ-rouen.fr

⁵Contact: helene.benghalem@gmail.com.

⁶Contact: emeline.limon@gmail.com

Abstract

This paper analyzes the consequences of the taxation of temporary jobs recently introduced in several European countries to induce firms to create more open-ended contracts and to increase the duration of jobs. The estimation of a job search and matching model on French data shows that the taxation of temporary contracts has consequences unexpected by policy makers: it reduces the mean duration of jobs. It also decreases job creation. This result holds even if the taxation is targeted at temporary contracts of short duration and is compensated by lower taxation of open-ended contracts and of temporary jobs of long duration. Accordingly, we find that the taxation of temporary contracts increases unemployment and reduces the welfare of unemployed workers.

Key words: Temporary jobs, Employment protection legislation, taxation.

JEL classification: J63, J64, J68.

1 Introduction

The spread of temporary jobs is an important concern in countries with stringent employment protection legislation. When open-ended contracts breaches are more costly, firms use more temporary contracts. The duration of temporary contracts is frequently very short. This spread of job instability induces frictional unemployment and a strong segmentation of labor markets between permanent and temporary jobs.¹ To limit job instability and labor market segmentation, several European countries have decided to tax temporary contracts.

Temporary contracts of short durations are especially targeted in France, Portugal and Spain, while all temporary contracts are taxed in Italy. France introduced in 2013 a tax equal to 3 percent of gross wages for temporary contracts shorter than one month, and to 1.5 percent for those from 1 to 3 months. If the temporary contract is transformed into an open-ended contract, the tax is refunded.² In 2014, Portugal introduced an adjustment of the rate of social contribution according to the type of labor contract, increasing the employers contribution by 4 percentage points (from 22.75 percent to 26.75 percent) for temporary contracts of durations shorter than 15 days. In Spain, unemployment insurance contributions are higher for temporary contracts than for permanent contracts since 1997.³ Since 2009, temporary contracts of short durations are particularly targeted. There is a supplementary employer social contribution, equal to 36 percent of gross wages, for temporary contracts of duration shorter than one week. The Italian reform enacted in 2012 introduced a tax on all temporary contracts equal to 1.4 percent of gross wages. The tax is refunded if temporary contracts are transformed into open-ended contracts. The amount of the refund is limited to the last six monthly payments of the tax.

These policies aim at inducing firms to use temporary jobs of longer duration and to turn temporary contracts into open-ended contracts. However, as far as we are aware, almost nothing is known about the consequences of such policies, which nonetheless have non trivial effects. This paper aims at contributing to fill this gap. It shows that the taxation of temporary contracts does not necessarily reduce job instability. Obviously, the taxation of contracts of short duration may induce employers to substitute contracts of longer duration for contract

¹See Bassanini and Garnero (2013), Boeri (2011), Bentolila et al. (2012), Sala et al. (2012).

²See: http://www.unedic.org/sites/default/files/ani_securisation_de_1_emploi_11012013.pdf

³6.7 percent instead of 5.5 percent for employers and 1.6 percent instead of 1.55 percent for employees. http://www.seg-social.es/Internet_1/Trabajadores/CotizacionRecaudaci10777/Basesytiposdecotiza36537/index.htm

of shorter duration and to transform temporary contracts into open-ended contracts if this allows them to avoid the tax. This effect is amplified if the tax is refunded when temporary contracts are transformed into open-ended contracts. The reduction of job instability can also be amplified if higher taxes on temporary contracts of short duration are compensated by lower taxes on temporary contracts of long duration and on open-ended contracts. But higher taxes have opposing effects on the duration of temporary contracts. For instance, it is unlikely that 7 day contracts are transformed into one month contracts in response to a tax increase on contracts shorter than one month, but it can be optimal to reduce the duration of contracts from 7 days to 6 days, because employers have incentives to reduce the length of temporary contracts when they are less profitable. Hence, in theory, higher taxes on temporary contracts do not necessarily reduce job instability. Their impact depends on the design of the tax scheme and on the empirical context.

To shed light on the consequences of taxation of temporary contracts, we provide a model that explains the impact of taxes on the distribution of spells of temporary and permanent jobs. This model shows that higher taxes on temporary contracts induce two effects. First, there is substitution of temporary contracts of longer duration for contracts of shorter duration when only contracts of very short duration are taxed (as in France, Portugal and Spain) and substitution of open-ended contracts for temporary contracts when all temporary contracts are taxed, as in the Italian system. Second, there is a drop in the duration of temporary contracts that are not worth lengthening to circumvent taxation. The model also accounts for the negative impact of taxation on job creation. The structural estimation of the model on French data allows us to run simulations to evaluate the impact of different tax systems on the distribution of employment spells, on unemployment and on the welfare of unemployed workers.

We find that the taxation of temporary contracts has a negative impact on the labor market. First, it reduces the mean duration of jobs. Hence, the taxation of temporary contracts does not achieve its main objective, which is to reduce labor turnover. Second, the tax decreases job creation. Accordingly, unemployment increases and the welfare of unemployed workers falls.

Our paper is related to at least two strands of the literature. First, the model relies on Cahuc, Charlot and Malherbet (2016), which explains the distribution of durations of temporary contracts and the choice between open-ended and temporary contracts. This model shows that

the use of temporary contracts induces an excess of job turnover leading to production losses. Our paper complements this analysis by estimating the structural parameters of the model, by introducing taxation of temporary contracts, and by running simulation exercises to evaluate the impact of different tax systems. One of its contributions consists in estimating a model that explains the shape of the distribution of job spells when there are temporary jobs instead of a very limited number of moments of this distribution, typically the mean or the median, as usually done in the literature. In particular, our model allows us to explain the large share of temporary contracts of very short durations, which is displayed on figure 1. This figure shows that about 50% of temporary contracts are shorter than one month in France. Usual models, relying on the standard version of the model of Mortensen and Pissarides (1994), are not able to account for this feature of temporary job creation. Another contribution, with respect to the literature devoted to the analysis of employment protection legislation, is to provide a much more complete picture of the consequences of regulations that change the relative cost of temporary and permanent jobs. Our approach allows us to evaluate the impact of such regulations on the distribution of employment spells and on the choice between permanent and temporary jobs. This is an improvement with respect to the current literature that does not explain in a unified framework the choice between temporary and permanent contracts, the duration of temporary contracts and their transformation into permanent contracts.⁴ Our approach is especially suited to evaluate different tax systems, targeted either at temporary contracts of short durations, like in the French system, or generalized to all temporary contracts, like in the Italian system.

The paper is organized as follows. The model is presented in section 2. Section 3 presents the data and the estimation of the benchmark model where the job arrival rate is exogenous. Section 4 is devoted to the empirical evaluation of the impact of different systems of taxation of temporary contracts. Section 5 extends the benchmark model to account for the reaction of the job arrival rate to the taxation of temporary contracts. Section 6 concludes.

⁴See, among others: Bentolila et al. (2012), Berson and Ferrari (2015), Berton and Garibaldi (2012), Blanchard and Landier (2002), Boeri and Garibaldi (2007), Cahuc and Postel-Vinay (2002), Costain, Jimeno and Thomas (2010), Macho-Stadler et al. (2014), Portugal and Varejão (2009), Sala, Silva and Toledo (2012), Smith (2007).

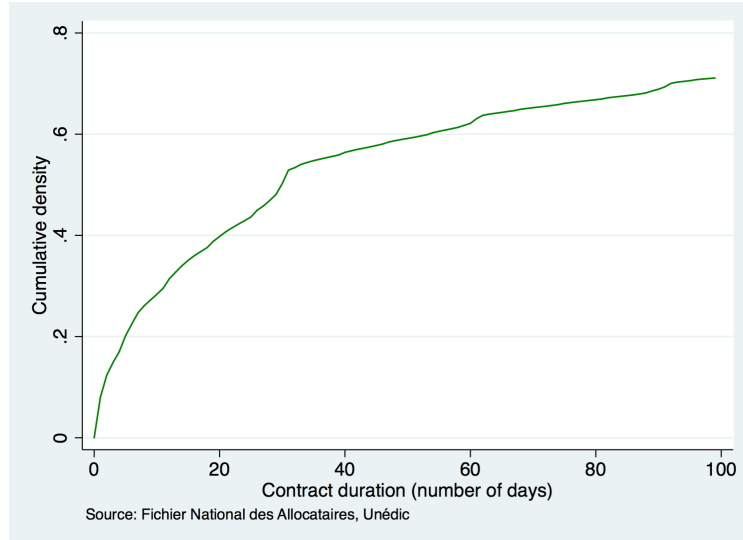


Figure 1: Cumulative density of temporary contracts durations in temporary jobs inflows in France over the period 2010-2012.

2 The model

This section outlines the economic environment in which we analyze the effects of the introduction of a tax on temporary contracts. We develop a job search model that builds on Cahuc, et al. (2016). In this framework, the choice (temporary or open-ended) and the duration of labor contracts are endogenous. Jobs can be either taxed or subsidized according to their type and duration. We first describe the framework before explaining how firms choose the type of contract and the duration of temporary jobs. Then, we define the labor market equilibrium.

2.1 The framework

2.1.1 Assumptions

Time is continuous and all agents discount the future at a common rate $r > 0$. There are two goods: labor, which is the sole input, and a numéraire good which is produced and consumed. Our analysis is focused on low wage workers, who are the most concerned by temporary jobs. In continental European countries, low wages are far from being competitive. They are set by wage floors at the national level and at the industry level. Accordingly, we assume that

the labor market is populated by a measure one of infinitely-lived unskilled workers paid the minimum wage w . Firms are competitive and create jobs to produce a numéraire output, using labor as sole input. All jobs produce the same quantity of output per unit of time, denoted by $y > 0$, but jobs differ by the rate at which they become unproductive, denoted by $\lambda > 0$. When a job is created, its type λ is randomly selected from $[\lambda_{\min}, +\infty)$, $\lambda_{\min} > 0$, according to a sampling distribution with cumulative distribution function G and density g . The distribution of λ has positive density over all its support and no mass point. Jobs and workers are brought together pairwise through a sequential, random and time consuming search process.

There are two types of contract: temporary and permanent. Permanent contracts stipulate the fixed minimum wage w and are open-ended: they do not stipulate any pre-determined duration. Permanent jobs can be terminated at any time at cost F . There is a (small) cost to write a contract, either temporary or permanent, which is denoted by $c > 0$.

Temporary contracts stipulate the wage w and a fixed duration. Temporary contracts are neither renegotiable nor renewable.⁵ The employer must pay the worker the wage stipulated in the contract until the date of termination, even if the job becomes unproductive before this date. At their date of termination, temporary jobs can be either destroyed at zero cost or transformed into permanent jobs.

Firms choose the type of contract that maximizes the value of the starting job. A temporary contract is chosen if it yields a higher value to the firm than a permanent contract. If a temporary contract is selected, the duration of the contract is chosen once for all in the starting contract because it is not permitted to renegotiate the contract.

Temporary contracts have to pay a tax. This tax, denoted by $\tau(\Delta) \geq 0$, can depend on the duration Δ of the temporary contract. Temporary contracts turned into open-ended contracts can get a refund, denoted by $\rho(\Delta)$, which can also depend on the duration of the temporary contract. The total amount of collected taxes is paid back to firms with a lump-sum subsidy to all jobs.

⁵The possibility to renew temporary contracts is analyzed in the working paper version of Cahuc et al. (2016), available as IZA discussion paper n°6365.

2.1.2 The value of permanent and temporary jobs

The value to a firm of starting permanent jobs with shock arrival rate λ , denoted by $J_p(\lambda)$, can be written as:

$$J_p(\lambda) = \int_0^\infty \left[\int_0^{\varkappa} (y - w) e^{-rt} dt - F e^{-r\varkappa} \right] \lambda e^{-\lambda\varkappa} d\varkappa - c. \quad (1)$$

The first term inside brackets, $\int_0^{\varkappa} (y - w) e^{-rt} dt$, stands for the discounted sum of expected profits, equal to the difference between y , the production, and w , the minimum wage, multiplied by the term e^{-rt} , which stands for the discount factor. Profits are expected until some random date \varkappa , at which the job becomes unproductive and is destroyed at cost F . The term $\lambda e^{-\lambda\varkappa}$ corresponds to the density of the Poisson process governing productivity shocks. The last term, c , denotes the cost to write the contract. The value $J_p(\lambda)$ can be written as:

$$J_p(\lambda) = \frac{y - w - \lambda F}{r + \lambda} - c. \quad (2)$$

By the same token, the value to a firm of starting temporary jobs with shock arrival rate λ and duration Δ , $J_t(\lambda, \Delta, \tau, \rho)$, can be written as:

$$J_t(\lambda, \Delta, \tau(\Delta), \rho(\Delta)) = \int_0^\Delta [y e^{-\lambda\varkappa} - w - \tau(\Delta)] e^{-r\varkappa} d\varkappa + \max [J_p(\lambda) + \rho(\Delta), 0] e^{-(r+\lambda)\Delta} - c. \quad (3)$$

The first term, $\int_0^\Delta [y e^{-\lambda\varkappa} - w - \tau(\Delta)] e^{-r\varkappa} d\varkappa$, stands for the discounted sum of expected profits over the duration of the job. In this expression, the level of production y is multiplied by the survival function $e^{-\lambda\varkappa}$ because the production drops to zero at rate λ . The wage w and the tax $\tau(\Delta)$ are not multiplied by the survival function because the employer has to keep and pay the employee until the date of termination of the contract. The second term, $\max [J_p(\lambda) + \rho(\Delta), 0] e^{-(r+\lambda)\Delta}$, is the present value of the option for the firm linked to the possibility of transforming the temporary job into a permanent job at the date of termination of the temporary contract, where $J_p(\lambda)$ is the value of the permanent job and $\rho(\Delta)$ stands for the refund of the tax paid on the temporary contracts. The present value of this option decreases with the duration of the contract because time is discounted at rate r and because the probability that the job is productive at the date of termination of the contract decreases with the spell of the contract. The last term is the cost to write the contract.

Let us now describe the optimal choice of the type of contract and of the duration of temporary contracts in the simplest case where there is no tax and no refund ($\tau(\Delta) = \rho(\Delta) = 0$). Then, we will analyze the consequences of the Italian and French tax systems.

2.2 The benchmark case without tax

In order to determine the choice between temporary and permanent contracts, we first need to define the value of temporary jobs at their optimal duration.

2.2.1 The optimal duration of temporary contracts

When $\tau(\Delta) = \rho(\Delta) = 0$, the optimal duration of temporary contracts is given by:

$$\Delta(\lambda) = \arg \max_{\Delta} J_t(\lambda, \Delta, 0, 0)$$

Using equation (3), it turns out that the optimal duration of a temporary contract on a job with shock arrival rate λ is defined by the following condition:⁶

$$\underbrace{ye^{-\lambda\Delta}}_{\text{marginal benefit}} = \underbrace{w + (r + \lambda) \max [J_p(\lambda), 0]}_{\text{marginal cost}} e^{-\lambda\Delta}. \quad (4)$$

In this expression, the left-hand side term, $ye^{-\lambda\Delta}$ stands for the marginal gain of an increase in the duration of the contract. This gain decreases with the duration of the contract because the survival probability of production opportunities decreases with the contract spell. It goes to zero when the duration goes to infinite. The right hand side corresponds to the marginal cost, which is equal to the sum of two terms. The first term, w , denotes the labor costs borne by the firm, and that must be paid, irrespective of the value of productivity, when the duration of the contract is increased. The second term, $(r + \lambda) \max [J_p(\lambda), 0] e^{-\lambda\Delta}$, is the option value linked to the possibility of transforming the temporary job into a permanent job. The marginal cost decreases with the duration of the job and has a strictly positive lower bound, equal to w .

The first order condition yields, together with equation (2), the optimal duration, as a function of λ , denoted by:

$$\Delta(\lambda) = \begin{cases} \frac{1}{\lambda} \log \left(\frac{w + \lambda F + (r + \lambda)c}{w} \right) & \text{if } \lambda \leq \lambda_p \\ \frac{1}{\lambda} \log \left(\frac{y}{w} \right) & \text{otherwise} \end{cases} \quad (5)$$

⁶The SOC is always fulfilled. Namely, it reads

$$-\lambda ye^{-\lambda\Delta} + \lambda e^{-\lambda\Delta} (r + \lambda) \max [J_p(\lambda), 0]$$

It is obviously negative when $\lambda > \lambda_p$, or equivalently, when $\max [J_p(\lambda), 0] = 0$, as it writes simply $-\lambda ye^{-\lambda\Delta} < 0$ in this case. When $J_p(\lambda) > 0$, the derivative of the first order condition with respect to Δ is $-\lambda ye^{-\lambda\Delta} + \lambda e^{-\lambda\Delta} (r + \lambda) J_p(\lambda)$ which is equal to (using (4)): $-\lambda w < 0$.



Figure 2: The relation between the shock arrival rate λ and the optimal duration of temporary contracts $\Delta(\lambda)$. Temporary jobs can be transformed into permanent jobs if $\lambda < \lambda_p$. Otherwise, they are destroyed at the end of the temporary contract.

where

$$\lambda_p = \{\lambda | J_p(\lambda) = 0\} \quad (6)$$

denotes the threshold value above which permanent jobs are no longer profitable. Equation (5) shows that function Δ decreases with λ ,⁷ with a kink at λ_p as shown on figure 2. When the shock arrival rate is smaller than λ_p , temporary jobs that have not been hit by a shock at duration $\Delta(\lambda_p)$ are transformed into permanent jobs. When the shock arrival rate is larger than λ_p , temporary jobs are destroyed when they reach their termination date. In other words, only temporary jobs with duration longer than $\Delta(\lambda_p)$ can be transformed into permanent jobs when they reach their termination date.

⁷It is easy to check that $\lim_{\lambda \rightarrow 0} \Delta(\lambda) = +\infty$ and $\lim_{\lambda \rightarrow \infty} \Delta(\lambda) = 0$.

2.2.2 The choice between temporary and permanent contract

Firms prefer permanent jobs to temporary jobs if and only if the value of starting permanent jobs is bigger than the value of starting temporary jobs, or more formally

$$J_p(\lambda) \geq J_t(\lambda) = \max_{\Delta} J_t(\lambda, \Delta, 0, 0).$$

Figure 3 displays the shape of the values of permanent and temporary jobs.⁸ It shows that permanent jobs are more profitable than temporary jobs if the shock arrival rate is smaller than the threshold value

$$\lambda_s = \{\lambda | J_p(\lambda) = J_t(\lambda)\}. \quad (7)$$

Accordingly, in that case, firms create permanent jobs. Otherwise, they create temporary jobs if the shock arrival rate is lower than

$$\lambda_t = \{\lambda | J_t(\lambda) = 0\}. \quad (8)$$

If a temporary job is created, its duration is equal to $\Delta(\lambda)$ and it can be transformed into a permanent job only if its duration is longer than $\Delta(\lambda_p)$.

2.3 The Italian system

In the Italian system, all temporary jobs pay the tax $\tau(\Delta) = \bar{\tau}$ independent of the duration of the contract and employers get a refund limited to the last 6 monthly payments of the tax. For the sake of simplicity, we assume that only temporary jobs of duration longer than 6 months are transformed into permanent jobs, meaning that the refund does not depend on the duration of the temporary jobs transformed into permanent jobs, i.e. $\rho(\Delta) = \bar{\rho} < \bar{\tau}\Delta$. The tax receipt is paid back to firms with a lump-sum subsidy to all jobs denoted by \bar{s} . It is easily checked that the optimal duration of temporary jobs is

$$\Delta^{\text{It}}(\lambda) = \begin{cases} \frac{1}{\lambda} \log \left(\frac{w - \bar{s} + \lambda F + (r + \lambda)(c - \bar{\rho})}{w - \bar{s} + \bar{\tau}} \right) & \text{if } \lambda \leq \lambda_p^{\text{It}} \\ \frac{1}{\lambda} \log \left(\frac{y}{w - \bar{s} + \bar{\tau}} \right) & \text{otherwise} \end{cases} \quad (9)$$

where $\lambda_p^{\text{It}} = \{\lambda | J_p(\lambda) + \bar{\rho} = 0\}$ is the threshold value of the shock arrival rate below which temporary jobs can be transformed into permanent jobs. From the definition of λ_p^{It} , it appears

⁸Formal proofs for the precise shape of these functions are given in appendix A.1.

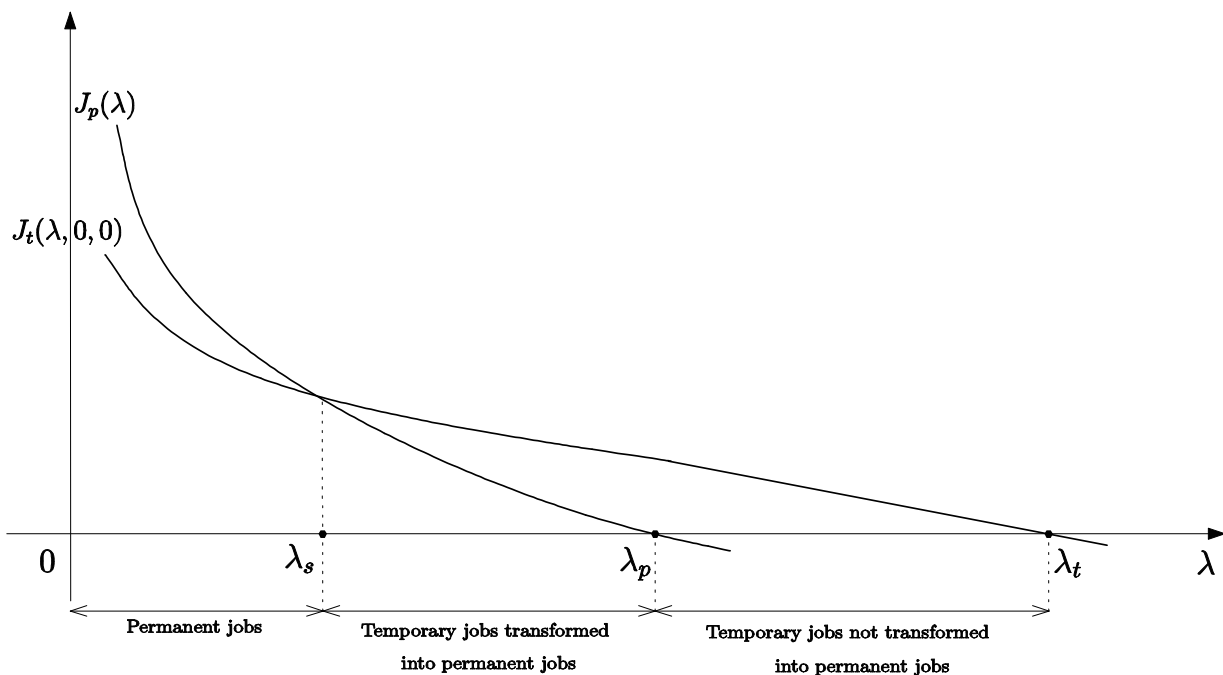


Figure 3: The relation between the shock arrival rate and the type of job creation.

that higher refunds $\bar{\rho}$ increase the share of temporary jobs transformed into permanent jobs.⁹ The refund also reduces the optimal duration of temporary contracts that can be transformed into permanent contracts. This effect contributes to increase the share of temporary jobs transformed into permanent jobs, since the probability to be hit by a shock increases with contract duration. The tax $\bar{\tau}$ reduces the duration of all temporary contracts. The tax also has a negative impact on the creation of temporary jobs since it reduces the threshold value of the shock arrival rate below which it is profitable to create temporary jobs, $\lambda_t^{\text{It}} = \{\lambda | J_t(\lambda, \bar{\tau}, \bar{\rho}) = 0\}$.

The tax and the refund modify the choice between temporary and permanent contracts. The condition to prefer permanent to temporary jobs is

$$J_p(\lambda) \geq J_t(\lambda, \bar{\tau}, \bar{\rho}) = \max_{\Delta} J_t(\lambda, \Delta, \bar{\tau}, \bar{\rho}).$$

The tax reduces the present value of starting temporary jobs since the refund $\bar{\rho}$ does not fully compensate the total expected amount of tax paid on temporary jobs. By reducing the relative profitability of temporary jobs, the Italian reform raises the number of creations of

⁹Since $J_p(\lambda)$ decreases with λ , the condition which defines λ_p^{It} , $J_p(\lambda_p^{\text{It}}) + \bar{\rho} = 0$ implies that $d\lambda_p^{\text{It}}/d\bar{\rho} > 0$.

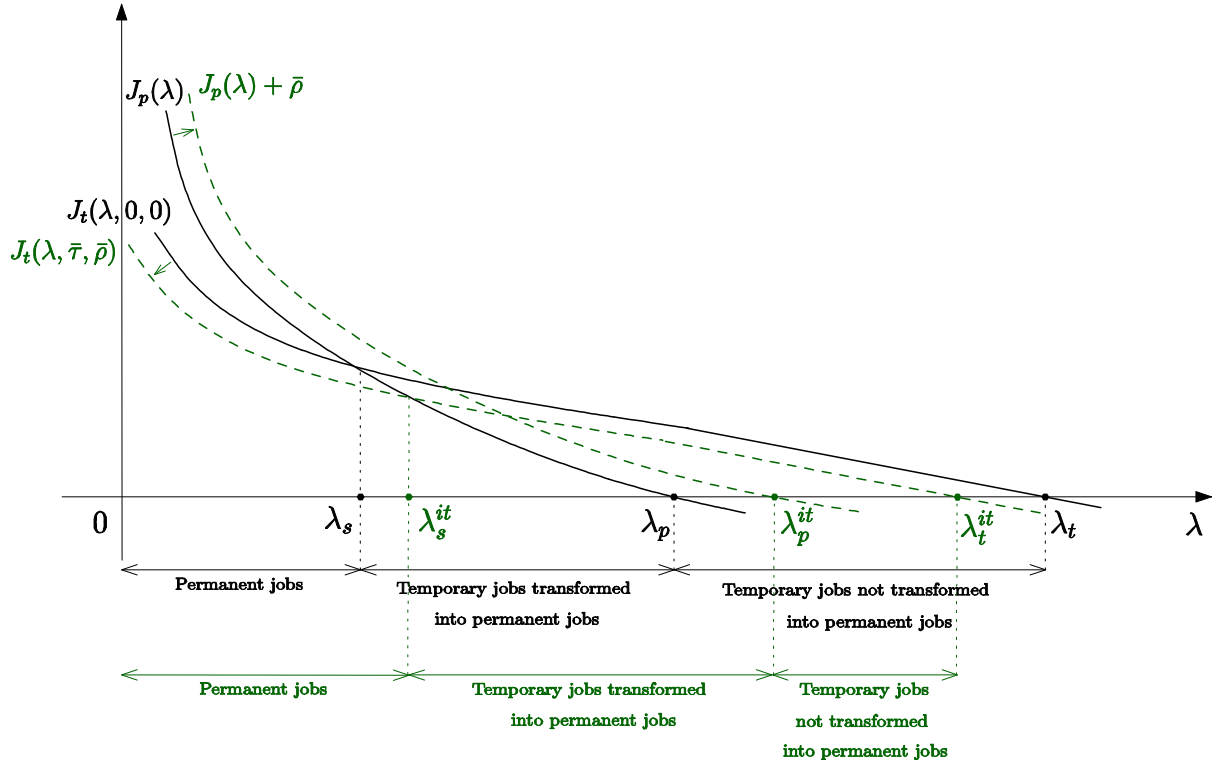


Figure 4: The relation between the shock arrival rate and the type of job creation in the Italian system.

permanent jobs ($\lambda_s^{\text{It}} = \{\lambda | J_p(\lambda) = J_t(\lambda, \bar{\tau}, \bar{\rho})\} > \lambda_s$) and the number of temporary jobs that are transformed into permanent jobs ($\lambda_p^{\text{It}} > \lambda_p$), as shown on figure 4 which displays the shape of functions $J_p(\lambda)$ and $J_t(\lambda, \bar{\tau}, \bar{\rho})$. These two effects contribute to decrease unemployment. However, the Italian reform decreases the total number of job creations since it lowers the threshold value of the shock arrival rate below which jobs are created ($\lambda_t^{\text{It}} < \lambda_t$). This is a direct consequence of the increase in labor costs induced by the tax, the refund being smaller than the total expected amount of taxes paid by firms. Another consequence of the increase in labor cost is the shorter duration of temporary jobs. These two effects contribute to increase unemployment.

All in all, the Italian reform has ambiguous effect on job stability, job market segmentation and unemployment. On one hand, it increases the number of permanent jobs, but on the other hand it reduces the duration of temporary jobs and prevents the creation of temporary jobs used to exploit production opportunities of very short durations (i.e. $\lambda > \lambda_t^{\text{It}}$).

2.4 The French system

The French system (including Portugal and Spain) targets the tax on temporary jobs of short durations. The tax $\bar{\tau}$ is paid for temporary contracts of duration shorter than $\bar{\Delta}$.¹⁰ The tax receipt is paid back to firms with a lump-sum subsidy to all jobs denoted by \bar{s} . There is a refund equal to the total amount of tax paid on temporary contracts transformed into open-ended contracts. For the sake of simplicity, we assume that $\bar{\Delta}$ is small enough to ensure that contracts of duration $\Delta \leq \bar{\Delta}$ subject to the tax are not transformed into permanent contracts. This assumption is relevant to describe the French system in which only contracts of durations shorter than 3 months are taxed.¹¹

This assumption implies that the duration of temporary contracts that can be transformed into permanent contracts is too long to be subject to the tax (i.e. longer than $\bar{\Delta}$), or, in other words, that firms do not pay taxes and do not get refund when they decide to transform temporary contracts into permanent contracts. In this context, the value to a firm of starting temporary jobs with shock arrival rate λ and duration Δ , is equal to

$$J_t(\lambda, \Delta, \bar{\tau}, 0) = \begin{cases} \int_0^{\Delta} (ye^{-\lambda z} - w + \bar{s}) e^{-rz} dz + \max [J_p(\lambda), 0] e^{-(r+\lambda)\Delta} - c & \text{if } \bar{\Delta} \leq \Delta \\ \int_0^{\Delta} (ye^{-\lambda z} - w - \bar{\tau} + \bar{s}) e^{-rz} dz - c & \text{if } \bar{\Delta} > \Delta. \end{cases} \quad (10)$$

The relation between the optimal duration of temporary contracts and the shock arrival rate is displayed on figure 5.¹² To understand the shape of the optimal duration, it is convenient to start from a low shock arrival rate and see how the duration changes as the arrival rate increases. When the shock arrival rate is sufficiently small, the optimal contract duration is longer than $\bar{\Delta}$, which implies that there is no tax to pay. In that case, the optimal duration of temporary contract is identical to that defined absent taxation, as defined by equation (5), except that the labor cost is equal to $w - \bar{s}$ instead of w . Therefore, the optimal contract duration is defined by:

$$\Delta^{\text{Fr}}(\lambda) = \begin{cases} \frac{1}{\lambda} \log \left(\frac{w - \bar{s} + \lambda F + (r + \lambda)c}{w - \bar{s}} \right) & \text{if } \lambda \leq \lambda_p^{\text{Fr}} \\ \frac{1}{\lambda} \log \left(\frac{y}{w - \bar{s}} \right) & \text{if } \lambda_p^{\text{Fr}} < \lambda \leq \bar{\lambda} \end{cases} \quad (11)$$

¹⁰ Actually, the French system comprises two thresholds. The tax amounts to 3 percent of the gross wage for contracts of duration shorter than one month and to 1.5 percent for contracts of duration from 1 to 3 months. We consider only one threshold for the sake of simplicity.

¹¹ This threshold is equal to 2 weeks in Portugal and 1 week in Spain.

¹² See appendix A.2 for a formal derivation of the optimal duration of temporary contracts in the French system.

where $\lambda_p^{\text{Fr}} = \{\lambda | J_p(\lambda) = 0\}$ is the threshold value of the shock arrival rate below which temporary jobs can be transformed into permanent jobs.¹³ This threshold value is larger than absent taxation ($\lambda_p^{\text{Fr}} > \lambda_p$) because the subsidy \bar{s} , which lowers the labor cost, increases the incentive to keep jobs at the termination date of temporary contracts. The fall of labor cost also increases the duration of temporary jobs. However, these effects are very small to the extent that the subsidy \bar{s} , which redistributes the tax receipt to all jobs, is very small in a context where only contracts of short durations are taxed.

Below the threshold $\bar{\lambda} = \{\lambda | \Delta^{\text{Fr}}(\lambda) = \bar{\Delta}\}$, the optimal contract duration is not directly influenced by the tax on temporary contracts: it is only affected by the lump-sum subsidy that slightly raises the duration of contracts. Now, if the shock arrival rate is higher than the threshold value $\bar{\lambda}$, the firm has to pay the tax τ if it chooses a duration lower than $\bar{\Delta}$. But it is not always profitable to do so. It can be more profitable to choose a duration equal to $\bar{\Delta}$, in order to avoid paying the tax. It is profitable to do so if the shock arrival rate is not too large, i.e. if $\lambda \leq \lambda_\tau = \{\lambda | \max_{\Delta} J_t(\lambda, \Delta, \bar{\tau}, 0) = J_t(\lambda, \bar{\Delta}, 0, 0)\}$. This implies that there is bunching at duration $\bar{\Delta}$, for all values of the shock arrival rate belonging to the interval $[\bar{\lambda}, \lambda_\tau]$, because it is worth lengthening contract duration to avoid the tax over this interval. If the shock arrival rate exceeds λ_τ , it becomes profitable to reduce the contract duration below $\bar{\Delta}$ and to pay the tax. In this situation, the tax reduces the optimal duration of temporary jobs as it increases the marginal cost of extending their duration.

In the French system, since only temporary contracts of very short duration are taxed, the choice between temporary and open-ended contracts is not directly impacted by the tax on temporary contracts. It is however influenced by the subsidy, which increases the relative value of permanent jobs. Accordingly, jobs start with open-ended contracts if the shock arrival rate is below the threshold λ_s^{Fr} which is larger than the threshold that prevails absent taxation, λ_s , displayed on figure 3, and with temporary contracts otherwise.

All in all, the French system increases the creation of permanent jobs and changes the duration of temporary contracts of short duration, with opposing effects. It decreases the very short durations and it increases the durations close to and shorter than the threshold duration below which temporary contracts are taxed.

¹³Here, $J_p(\lambda) = \frac{y-w+\bar{s}-\lambda F}{r+\lambda} - c$. We keep the same notation as in the case without taxation for the sake of simplicity.

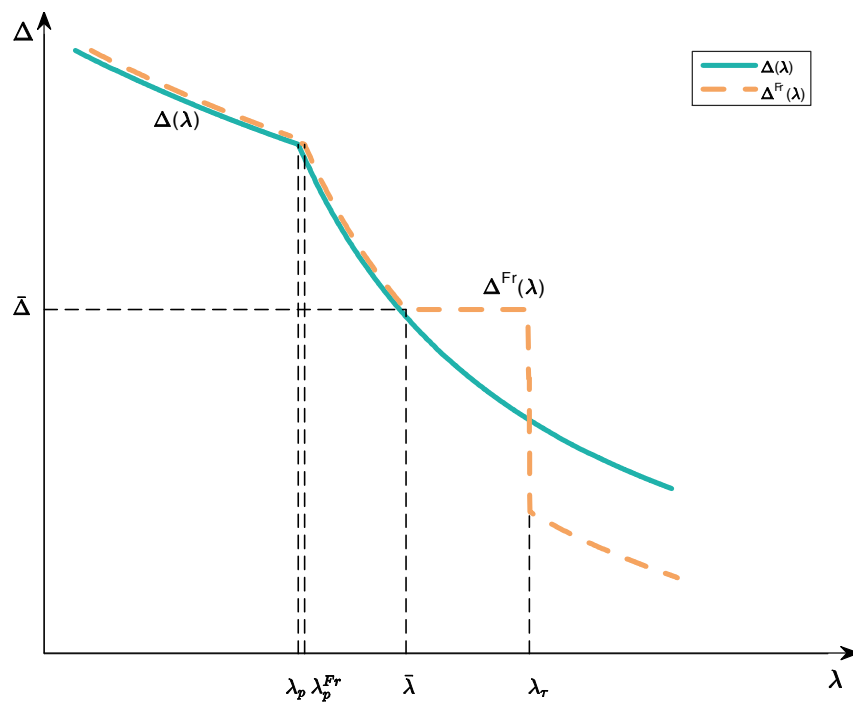


Figure 5: The relation between the shock arrival rate λ and the optimal duration of temporary jobs in the French system (dotted line) and in the system without tax (continuous line).

2.5 Unemployment and welfare

Each unemployed worker is matched with firms that allow him to get job opportunities at rate α . Since only jobs with productivity shock arrival rate below the threshold value λ_t are created, the job finding rate is $\alpha G(\lambda_t)$. The job finding rate, together with the equilibrium values of λ_p, λ_s and λ_t defined by equations (6), (7), (8) determines the equilibrium distribution of job durations and the equilibrium unemployment rate computed from the equality of unemployment inflows and outflows. Absent taxation of temporary contracts, the equilibrium unemployment rate is defined by:¹⁴

$$u = \frac{1}{1 + \alpha \left[\int_{\lambda_{\min}}^{\lambda_s} \frac{1}{\lambda} g(\lambda) d\lambda + \int_{\lambda_s}^{\lambda_p} g(\lambda) \left[\frac{e^{-\lambda\Delta(\lambda)}}{\lambda} + \Delta(\lambda) \right] d\lambda + \int_{\lambda_p}^{\lambda_t} g(\lambda) \Delta(\lambda) d\lambda \right]}. \quad (12)$$

It is also possible to compute the discounted expected utilities of unemployed workers and of workers on temporary jobs and on permanent jobs. Let us assume that workers have no access to financial markets and that production is non storable so that the flow of consumption is equal to the flow of income. Let us denote by $v(\cdot)$, $v'(\cdot) > 0$, $v''(\cdot) \leq 0$ the instantaneous utility function, which depends on instantaneous income. If b denotes unemployment benefits, the discounted expected utilities of unemployed workers W_u , of employees on type- λ starting temporary jobs, $W_t(\lambda)$, and on type- λ permanent jobs, $W_p(\lambda)$, satisfy

$$\begin{aligned} rW_u &= v(b) + \alpha \left[\int_{\lambda_{\min}}^{\lambda_s} [W_p(\lambda) - W_u] dG(\lambda) + \int_{\lambda_s}^{\lambda_t} [W_t(\lambda) - W_u] dG(\lambda) \right] \\ rW_p(\lambda) &= v(w) + \lambda [W_u - W_p(\lambda)] \\ W_t(\lambda) &= \begin{cases} \int_0^{\Delta(\lambda)} v(w) e^{-rt} dt + e^{-r\Delta(\lambda)} W_u & \text{if } \lambda_t \geq \lambda > \lambda_p \\ \int_0^{\Delta(\lambda)} v(w) e^{-rt} dt + e^{-(r+\lambda)\Delta(\lambda)} W_p(\lambda) + [1 - e^{-\lambda\Delta(\lambda)}] e^{-r\Delta(\lambda)} W_u & \text{if } \lambda_p \geq \lambda > \lambda_s \end{cases} \end{aligned}$$

3 Data and Estimation

3.1 Data

Information about employment spells comes from administrative records of the public employment agency (Pôle Emploi and Unédic). These records comprise information on the employment spells, on the type of contract, the wage, the number of hours worked and on several characteristics of firms and workers. These records cover all the contracts of the past work experience

¹⁴See appendix A.3.

of individuals registered to the public employment agency. They do not cover all the universe of labor contracts since individuals who never registered to the public employment agency are not covered. Comparison of these data with other sources that register all hiring intentions shows that the number of temporary jobs registered by the public employment agency covers about 70 percent of hiring intentions and evolves as that of the number of hiring intentions (Benghalem, 2016).

A natural strategy to analyze the impact of the taxation of temporary contracts is to look at the changes in the distribution of contract durations around the 1 month and the 3 month thresholds before and after the implementation of the taxation of temporary jobs, on 1 July 2013. Unfortunately, the actual tax implemented in France was so little effective that it does not allow us to proceed with this strategy. Many industries, professions and types of contract were exempted from the tax. For instance, temporary contracts used to replace absent workers and seasonal jobs were exempted. These exemptions created many leeways to avoid taxation. All in all, the amount of collected taxes has been very low.¹⁵ As a consequence, available data do not allow us to detect a change in the distribution of contract durations around the 1 month and 3 month thresholds before and after 1 July 2013, as shown by figure 6 which displays the distributions of contract durations from 1 July 2012 to 30 June 2013 and from 1 July 2013 to 30 June 2014 for the professions, types of contracts and industries subject to the tax. Nevertheless, our model allows us to evaluate the potential impact of the tax. The model is estimated over the period from January 2010 to June 2013, before the implementation of the taxation of temporary jobs.

3.2 Estimation

The block recursivity of the model allows us to proceed to the estimation of its parameters step by step. The first block of the model determines the distribution of contract spells, which is defined by equations (5), (6) (7), (8) and by the distribution G of the arrival rates of productivity shocks, which is assumed to be a Weibull distribution.¹⁶ This implies that the

¹⁵The annual receipt of the taxation of temporary jobs is about 70 millions euros. The total receipt for the contributions to unemployment insurance is about 30 billions euros. The amount of collected taxes represents only 1.5 percent of the wage for contracts of duration shorter than one month in eligible professions and industries (instead of 3 percent in principle) and 0.7 percent for contracts of duration from 1 to 3 months (instead of 1.5 percent). Accordingly, the change in tax at the 1 month and 3 months thresholds are very small.

¹⁶Estimates with a generalized Gamma distribution converges to a Weibull distribution.

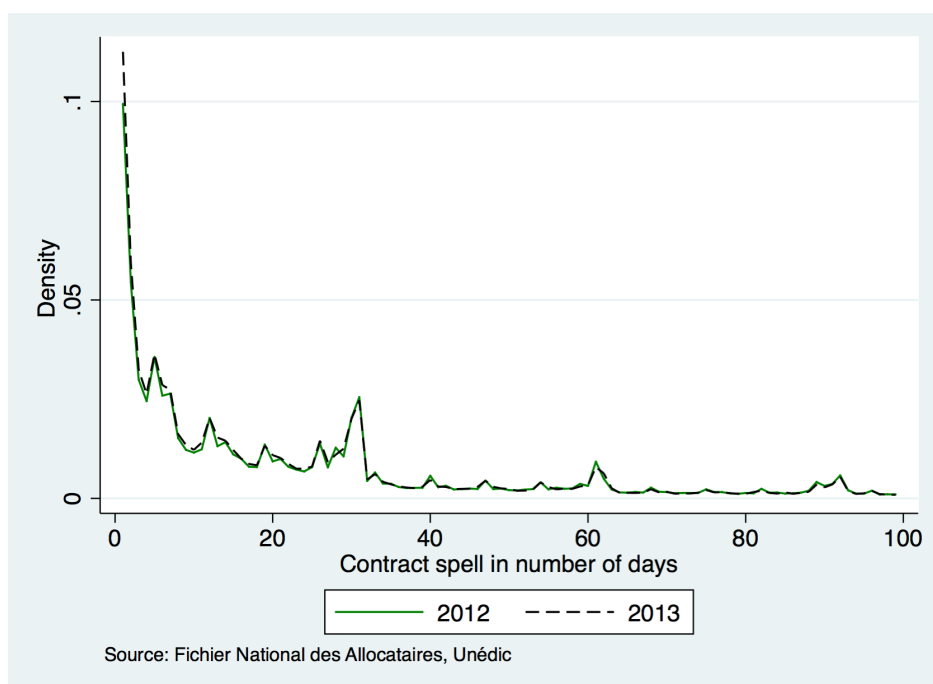


Figure 6: Density of durations of temporary contracts among all temporary contracts for professions, industries and types of contracts subject to the tax implemented on 1 July 2013. 2012: contracts starting from 1 July 2012 to 30 June 2013; 2013: contracts starting from 1 July 2013 to 30 June 2014.

distribution of contract spells is entirely defined by seven parameters: the discount rate r , the productivity y , the wage w , the scale μ , and the shape σ , of the Weibull distribution, the firing costs F and the costs of writing contracts c . The discount rate is set to 0.05 on an annual basis and the wage, which is exogenous in the model, is normalized to one. The five remaining parameters are estimated with the Generalized Method of Moments. The S moments are the shares of contracts of spell equal to $(1, 2, \dots, S)$ days. Let us denote by $p(s|\theta)$ the proportion of contracts of spell equal to s days predicted by the model conditional on the vector of parameters $\boldsymbol{\theta} = (\mu, \sigma, F, c, y)$ and denote by $p(s)$ the empirical proportion of contracts of spell equal to s days. The GMM estimator $\hat{\boldsymbol{\theta}}$ is defined by the following quadratic form

$$\hat{\boldsymbol{\theta}} = \arg \min_{\boldsymbol{\theta}} [\mathbf{p} - \mathbf{p}(\boldsymbol{\theta})] \hat{\Omega}^{-1} [\mathbf{p} - \mathbf{p}(\boldsymbol{\theta})]'$$

where $\mathbf{p} = (p(1), \dots, p(S))$, $\mathbf{p}(\boldsymbol{\theta}) = (p(1|\boldsymbol{\theta}), \dots, p(S|\boldsymbol{\theta}))$ and $\hat{\Omega}^{-1}$ is a symmetric and positive definite efficient weighting matrix.¹⁷ In the benchmark estimates, the vector $(1, \dots, S)$ is equal to job spells from 1 to 45 days. The results, presented in Table 1, are empirically relevant: we find that the wage amounts to 71% of productivity. The firing costs are equal to about two monthly wages (64 days) and the cost to write a contract represents about 0.08% of the daily wage. The fit between the empirical density of the contract durations and that predicted by the model, represented on Figure 7, is good. This visual impression is confirmed by the Hansen over-identification test as shown in Table 1.

Once the values of the parameters μ, σ, y, F, c are estimated, the value of the arrival rate of job opportunities α is chosen to match the unemployment rate of unskilled workers, equal to 13.5%.

4 Empirical evaluation

This section is devoted to the analysis of the impact of the taxation of temporary contracts in the French and in the Italian system.

¹⁷The estimation procedure is detailed in Appendix A.4.

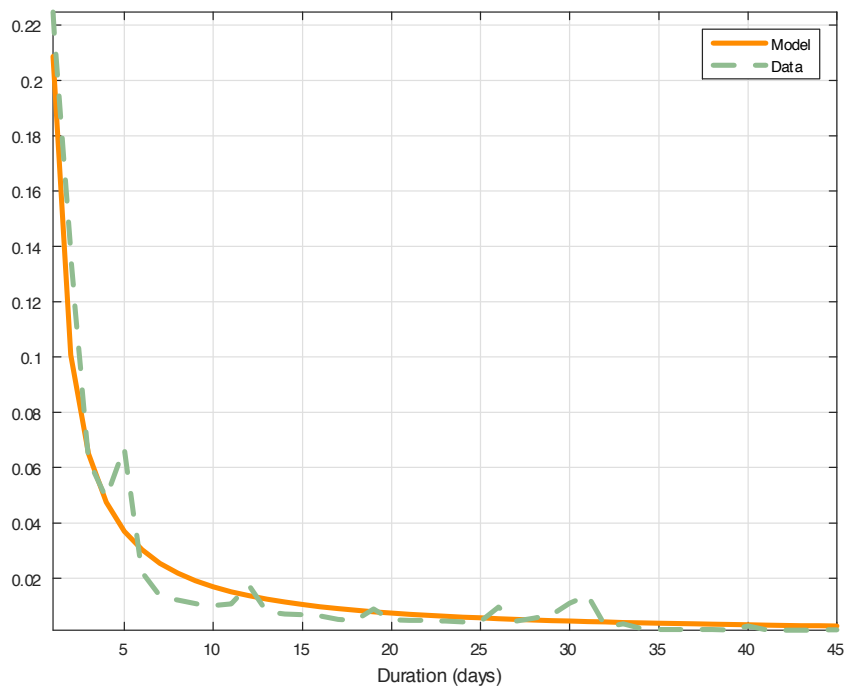


Figure 7: Empirical and estimated distributions of temporary contract durations in the flow of entries.

4.1 The French system

In France, Portugal and Spain, taxation of temporary contracts targets contracts of short duration. Figure 8 displays the consequences of a tax of 1.5% of the labor cost on temporary contracts of duration shorter than one month on the distribution of contract spells in the flow of entries into employment. The figure shows that there are no contracts between 23 and 30 days after the introduction of the tax because it is more profitable to use contracts of duration longer than one month to avoid taxation. There is bunching just above one month. The bunching increases the duration of contracts because contracts of durations between 23 days and one month are lengthened. However, the contracts below 23 days are shortened as shown in table 2 which reports the impact of the tax on unemployment, welfare and the duration of temporary contracts. The last column of table 2 shows that the tax decreases the mean duration of temporary contracts. This result is striking to the extent that the aim of the taxation of temporary jobs is to decrease job turnover. Our evaluation suggests that the policy has the opposite effect. The tax also induces a fall in λ_t , the threshold value of the shock arrival rate λ below which jobs are created. This reduces the exit rate from unemployment, equal to $\alpha G(\lambda_t)$.

All in all, the drop of the exit rate from unemployment and the decrease in the mean duration of temporary jobs imply that unemployment increases. However, the effect is small when the tax is targeted to contracts of short durations, as in France (maximum 3 months), Portugal (2 weeks) and Spain (1 week). A tax on contracts of duration shorter than one month equal to 10% of the labor cost raises unemployment by 0.004 percentage points.

Since the tax decreases the job finding rate and increases job turnover, its impact on the discounted expected utility of unemployed workers is negative. Here too, the effect is small. A tax on contracts of duration shorter than one month equal to 10% of the labor cost induces a drop of welfare of unemployed workers equivalent to a decrease in the unemployment benefit replacement ratio of 0.02%.

4.2 The Italian system

In the Italian system, all temporary contracts are taxed. The tax is refunded if temporary contracts are transformed into open-ended contracts, but the amount of the refund is limited

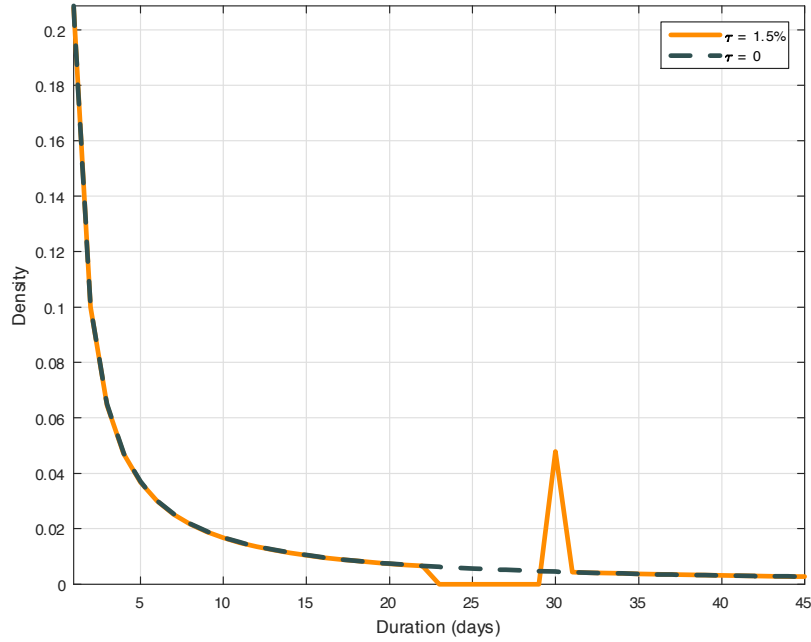


Figure 8: Impact of taxation of temporary contracts of duration shorter than one month on the distribution of the duration of contracts.

to the last six monthly payments of the tax. Since all temporary contracts are taxed, the tax has a stronger impact than in the French system, where only contracts of short duration are taxed. Table 3 shows that the duration of temporary contracts of short duration decreases as in the French system. The duration of temporary contracts of long spell (above one month in our framework), which is slightly increased thanks to the lump-sum subsidy \bar{s} in the French system, decreases in the Italian system. This contributes to amplify the negative effects of the tax on job stability. But the refund of the tax for temporary contracts transformed into permanent contracts induces more transformation of temporary contracts into permanent contracts. This counteracts the shortening of temporary contracts. Nevertheless, table 3 shows that the tax increases unemployment more in the Italian system than in the French system, merely because more temporary contracts are taxed in the Italian system. Unemployment increases by 0.03 percentage points when the tax equals 10%, an amount 9 times larger than in the French system. The fall of welfare of unemployed workers is also about 9 times larger in the Italian than in the French system.

5 Extensions

Until now, it has been assumed that the arrival rate of job offers was exogenous, equal to α . In this section, the arrival rate of job offers is made endogenous to account for the potential impact of the tax on temporary contracts on job creation.

5.1 Labor market equilibrium

In order to account for job creation, it is assumed that there is free entry on the labor market. Firms must pay the fixed cost $\kappa > 0$ to post a vacant job. κ is a sunk cost. Unemployed workers and job vacancies are brought together through a constant returns to scale matching technology which implies that vacant jobs are filled at rate $q(\theta)$, $q'(\theta) < 0$, where θ denotes the labor market tightness, equal to the ratio of the number of job vacancies over unemployment. Once matches are created, firms draw production opportunities from the sampling distribution $G(\lambda)$ of arrival rates of productivity shocks. The distribution of λ has positive density over all its support and no mass point. As shown in section 2.2, jobs are created only if the shock arrival rate is lower than the threshold λ_t . In this case, a temporary job is created if the shock arrival rate is bigger than λ_s and a permanent job is created otherwise. Thus, the value of a vacant job satisfies

$$rV = q(\theta) \left[\int_{\lambda_{\min}}^{\lambda_s} J_p(\lambda) dG(\lambda) + \int_{\lambda_s}^{\lambda_t} \max_{\Delta} J_t(\lambda, \Delta, 0, 0) dG(\lambda) - V \right], \quad (13)$$

where the value of $J_p(\lambda)$ is defined by equation (2), that of $\max_{\Delta} J_t(\lambda, \Delta, 0, 0)$ by equations (3) and (5). The free entry condition $V = \kappa$, can be written

$$\kappa = \frac{q(\theta)}{r + q(\theta)} \left[\int_{\lambda_{\min}}^{\lambda_s} J_p(\lambda) dG(\lambda) + \int_{\lambda_s}^{\lambda_t} \max_{\Delta} J_t(\lambda, \Delta, 0, 0) dG(\lambda) \right]. \quad (14)$$

The equilibrium distribution of job spells and the unemployment rate are defined as in the benchmark model, except that the variable $\theta q(\theta)$ is substituted for α , where θ is defined by equation (14). At this stage, thanks to the block recursivity of the model, we already know the empirical values of the parameters of the Weibull distribution (μ, σ) , of the productivity y (the wage w is normalized to 1), of the firing cost F and of the cost of writing contracts c . We need to determine the empirical values of the parameters of the matching function and of κ the cost of posting job vacancies to define the equilibrium value of the labor market tightness.

5.2 Estimation and calibration

We assume that the matching function is Cobb-Douglas and homogeneous of degree one, which implies that the number of hires, H , is defined by the expression $H = mU^{1-\eta}V^\eta$, $\eta \in (0, 1)$, $m > 0$, where U stands for the number of unemployed workers and V for the number of vacant jobs. Therefore, the exit rate from unemployment, H/U , can be written $m\theta^\eta$, with $\theta = V/U$. To evaluate the parameters m and η of the matching function we estimate the logarithm of the job finding rate:

$$\log(H/U) = \eta \log \theta + \nu$$

where $\nu = \log m$.

Data on unemployment and job vacancies for low skilled workers come from the French employment agency (*Pôle emploi*). Firms can post job vacancies at *Pôle emploi*. This is a free service and *Pôle emploi* estimates that they deal with almost 50% of the total of French vacancies. These data allow us to compute the labor market tightness, as the ratio of the number of job vacancies posted at the employment agency over the number of unemployed workers registered at the employment agency, at the commuting zone level for blue collars and low skilled white collars for each year from 2009 to 2011. There are 348 commuting zones. Data on hires of blue collars and low skilled white collars at the commuting zone level over the same period come from two data sets provided by the French Ministry of labor. The DMMO register (Déclaration Mensuelle de Mouvements de main d'Oeuvre), which describes establishments job flows (entries, exits, jobs created and lost, etc.) by type of contract, gender, age, occupational category. This is an administrative register which is mandatorily filled by all establishments with more than 50 employees. Information for establishments with less than 50 employees relies on the EMMO survey (Enquête sur les Mouvements de Main d'Oeuvre), which is a quarterly survey providing the same information as the DMMO register.

We use variations across commuting zones over time to identify the parameter η . We measure the tightness θ_{jt} and unemployment (U_{jt}) at the commuting zone level j at date t from the employment agency data and the hires (H_{jt}) from the establishment data. Let us denote by f_{jt} the annual job finding rate (H_{jt}/U_{jt}). We estimate the following equation

$$\log f_{jt} = a_1 \log \theta_{jt} + \sum_t b_t 1[\text{date} = t] + c_j + \nu_{jt} \quad (15)$$

where j is one of the 348 commuting zones and the date t varies from 2009 to 2011. The

estimation controls for date dummies and commuting zones fixed effects (c_j). The equation is estimated by standard (within) OLS regression, estimating equation (15) in first difference to eliminate the commuting zone fixed effect. However, the OLS estimates are exposed to an endogeneity bias arising from the search behavior of agents on either side of the market. For instance, improvements in the matching technology parameter m can raise the labor market tightness θ and the hiring rate (see Borowczik et al. 2013). This implies a potential correlation between the residuals of the OLS estimation and the labor market tightness which can bias downwards the OLS estimate of the coefficient of the labor market tightness. To address this issue, we proceed to IV estimation following the approach of Bartik (1993). The variation across years of the labor market tightness in commuting zone j is instrumented by the national rates of growth of the number of entries into employment across industries using commuting zone industry entries shares as weights to produce a measure of local labor demand shifts unrelated to changes in local matching technology and local labor supply. More formally, the instrument in commuting zone j in period t is $z_{jt} = \sum_i s_{jit-1} \text{dln } E_{it}$ where s_{jit-1} denotes the share of entries of industry i in commuting zone j in year $t - 1$ and $\text{dln } E_{it} = \text{dln } \sum_j e_{jit}$, where e_{jit} denotes the growth rate of the number of entries in commuting zone j in industry i in year t . The lagged values of z_{jt} are also used as instruments. These instruments are strongly correlated with the labor market tightness, as shown by table 4. Table 4 shows the estimates of the coefficient a_1 using OLS in column 1 and IV in column 2. Both estimates are highly significant. However, the OLS estimate is significantly lower than the IV estimate as expected. Taking the IV estimation as our preferred estimate, η , the elasticity of the matching function with respect to the number of vacancies, amounts to 0.50.

The estimation of the matching function allowed us to estimate the value of the elasticity of the matching function with respect to the number of vacancies. One needs to define the value of 2 more parameters, κ the cost of posting job vacancies, and m , the parameter of the matching function, to be able to define the equilibrium value of the labor market tightness, defined by the free entry condition (14). The values of these two parameters are chosen to match the unemployment rate, equal to 13.5%, and the elasticity of employment with respect to the wage, assumed equal to 1, which is the relevant target for low skilled workers (Hamermesh, 2014).

5.3 Results

The reaction of labor market tightness amplifies the negative impact of the taxation of temporary contracts on the labor market as shown by table 5 for the French system, where only temporary contracts of short duration are taxed, and by table 6 for the Italian system, where all temporary contracts are taxed. The impact of the tax on the duration of contracts is not reported in these tables because it is almost identical to the case where the arrival rate of job offers is exogenous, displayed in tables 2 and 3.

In the benchmark case for the French system, reported in columns 3 and 4 of table 5, where the elasticity of the matching function η is equal to 0.5, the impact on unemployment of the tax equal to 10% of the labor cost is about 15 times larger (0.06 percentage points instead of 0.004) than when the job arrival rate is exogenous. In order to gauge the robustness of this result, columns 1 and 2 of table 5 present the results when the elasticity of the matching function equals 0.39 instead of 0.5, which implies that the wage elasticity of employment is equal to 0.39 instead of 1. Although the wage elasticity of employment is much lower – actually a lower bound for the wage elasticity of employment of low skilled workers –, the reaction of labor market tightness still considerably amplifies the impact of the tax on unemployment, which is 7 times larger (0.03 percentage points instead of 0.004) than when the job arrival rate is constant. Columns 5 and 6 of table 5 show that the unemployment rate increases by 0.15 percentage points when the elasticity of the matching function is equal to 0.6, corresponding to the wage elasticity of employment equal to 2.73.

Comparison of welfare of unemployed workers in tables 2 and 5 shows that the negative impact of the tax of 10% of the labor cost on welfare is about 7 times larger than when the job arrival rate is exogenous in the benchmark case where the wage elasticity of employment equals 1. This ratio falls to 3 when the wage elasticity of employment equals 0.4 and climbs to 15 when the wage elasticity of employment equals 2.73.

The analysis of the Italian system relying on the comparison on tables 3 and 6 leads to the same conclusion: the reaction of labor market tightness considerably amplifies the negative impact of the taxation of temporary contracts on employment and welfare.

6 Conclusion

Our analysis suggests that the taxation of temporary contracts is not an appropriate policy to induce firms to create more stable jobs on a typical Continental European labor market that features stringent protection of permanent jobs and downward wage rigidity. The taxation of temporary contracts shortens the average job duration, raises unemployment and reduces the welfare of unemployed workers. This conclusion holds even if the taxation is targeted at temporary contracts of short duration and is compensated by lower taxation of open-ended contracts and of temporary jobs of long duration. All in all, the taxation of temporary contracts deteriorates labor market efficiency.

This result suggests that other policies should be implemented to fight against the strong segmentation of European labor markets between permanent and temporary jobs. Reducing dismissal costs for open-ended contracts, introducing an experience-rated system as in the US or layoff taxes that steadily increase with seniority is likely more appropriate.¹⁸

¹⁸Blanchard and Tirole (2008), Dolado et al. (2016), Garcia-Perez and Osuna (2014), L'Haridon and Malherbet (2009).

References

- [1] Bartik, T. J., 1993, Who benefits from local job growth: migrants or the original residents?, *Regional studies*, 27(4), 297–311.
- [2] Bassanini, A., and A. Garnero, 2013, Dismissal Protection and Worker Flows in OECD Countries: Evidence from Cross-country/Cross-industry Data, *Labour Economics*, 21(1), 25–41.
- [3] Benghalem, H. 2016, La majorité des embauches en contrat court se font chez un ancien employeur, *Eclairages, Etudes et Analyses*, n°14, UNEDIC.
- [4] Bentolila, S., Cahuc, P., Dolado, J. and T. Le Barbanchon, 2012, Two-Tier Labor Markets in a Deep Recession: France vs. Spain, *Economic Journal*, 122, , F155–F187.
- [5] Berson, C., Ferrari, N. 2015, Financial incentives and labour market duality, *Labour Economics*, vol 37, pp. 77–92.
- [6] Berton, F. and Garibaldi, P., 2012, Workers and Firm Sorting into Temporary Jobs, *Economic Journal*, 122, F125–F154.
- [7] Blanchard, O. J. and A. Landier, 2002, The Perverse Effects of Partial Labor Market Reform: Fixed Duration Contracts in France, *Economic Journal*, 112, 214–244.
- [8] Blanchard, O. J. and Tirole, J. 2008, The joint design of unemployment insurance and employment protection: A first pass. *Journal of the European Economic Association*, 6(1):45–77.
- [9] Boeri, T. 2011, Institutional Reforms in European Labour Markets, *Handbook of Labor Economics*, Elsevier, vol. 4B, chap 13, 1174–1235.
- [10] Boeri T. and P. Garibaldi , 2007, Two Tier Reforms of Employment Protection Legislation. A Honeymoon Effect?, *Economic Journal*, 117, F357–F385.
- [11] Borowczyk-Martins, D., G. Jolivet, and F. Postel-Vinay, 2013. Accounting For Endogeneity in Matching Function Estimation, *Review of Economic Dynamics*, 16, 440–51.
- [12] Cahuc, P., O. Charlot and F. Malherbet, 2016, Explaining the Spread of Temporary Jobs and Its Impact on Labor Turnover, *International Economic Review*, forthcoming.
- [13] Cahuc, P. and F. Postel-Vinay, 2002, Temporary Jobs, Employment Protection and Labor Market Performance, *Labour Economics*, 9, 63–91.
- [14] Costain, J., Jimeno, J. and C. Thomas, 2010, Employment Fluctuations in a Dual Labor Market, Banco d’Espana, Working paper 1013.

- [15] Dolado, J., Lalé, E., Siassi, N., 2016, From Dual to Unified Employment Protection: Transition and Steady State, IZA working paper No. 9953.
- [16] Garcia Perez, I., and Osuna, V., 2014, Dual Labour Markets and the Tenure Distribution: Reducing Severance Pay or Introducing a Single Contract, *Labour Economics*, 29, 1-13.
- [17] Hamermesh, D., 2014., Do labor costs affect companies' demand for labor? IZA World of Labor 2014: 3
- [18] L'Haridon, O. and F. Malherbet, 2009, Employment Protection Reform in Search Economies, *European Economic Review*, 53, 255-273.
- [19] Macho-Stadler, I., Pérez-Castrillo, D. and N. Porteiro, 2014, Coexistence of Long-term and Short-term Contracts, *Games and Economic Behavior* 86, 145-164.
- [20] Mortensen, D. T. and C. A. Pissarides, 1994, Job Creation and Job Destruction in the Theory of Unemployment, *Review of Economic Studies*, 61, 397-415.
- [21] Portugal, P. and J. Varejão, 2009, Why Do Firms Use Fixed-Term Contracts? IZA discussion paper 4380.
- [22] Sala, H., J. I. Silva, and M. E. Toledo, 2012, Flexibility at the Margin and Labor Market Volatility in OECD Countries, *Scandinavian Journal of Economics*, 114, 991-1017.
- [23] Smith, E., 2007, Limited Duration Employment, *Review of Economic Dynamics*, 10, 444-471.

A Appendix

A.1 Properties of the values of permanent and temporary jobs

A.1.1 Properties of $J_p(\lambda)$

The function:

$$J_p(\lambda) = \frac{y - w - \lambda F}{r + \lambda} - c \quad (\text{A1})$$

is continuous. It is decreasing in λ , as $J'_p(\lambda) = -\frac{y-w+rF}{(r+\lambda)^2} \leq 0$. It decreases from $\lim_{\lambda \rightarrow 0} J_p(\lambda) = \frac{y-w}{r} - c \geq 0$ to $\lim_{\lambda \rightarrow +\infty} J_p(\lambda) = -c - F \leq 0$. Thus, there exists a unique threshold

$$\lambda_p = \frac{y - w - rc}{c + F} \quad (\text{A2})$$

such that $J_p(\lambda_p) = 0$ and $J_p(\lambda) \geq 0$ iff $\lambda \leq \lambda_p$, as indicated in text.

A.1.2 Properties of $J_t(\lambda) = \max_{\Delta} J_t(\lambda, \Delta)$

The value of a temporary job is:

$$J_t(\lambda) = y \left(\frac{1 - e^{-(r+\lambda)\Delta(\lambda)}}{r + \lambda} \right) - \frac{w}{r} \left(1 - e^{-r\Delta(\lambda)} \right) + \max [J_p(\lambda), 0] e^{-(r+\lambda)\Delta(\lambda)} - c. \quad (\text{A3})$$

Function $J_t(\lambda)$ is continuous over $[0, +\infty[$ and has a kink at $\lambda = \lambda_p$. It decreases from $\lim_{\lambda \rightarrow 0} J_t(\lambda) = \frac{y-w}{r} - c \geq 0$ to $\lim_{\lambda \rightarrow \infty} J_t(\lambda) = -c \leq 0$. Let us prove that $J_t(\lambda)$ is decreasing in λ .

- When $\lambda \geq \lambda_p$ we have

$$J_t(\lambda) = y \left(\frac{1 - e^{-(r+\lambda)\Delta(\lambda)}}{r + \lambda} \right) - \frac{w}{r} \left(1 - e^{-r\Delta(\lambda)} \right) - c$$

Keeping in mind that the envelope theorem implies that $\partial J_t / \partial \Delta = 0$, we have

$$J'_t(\lambda) = y \frac{\Delta(\lambda) (r + \lambda) e^{-(r+\lambda)\Delta(\lambda)} - 1 + e^{-(r+\lambda)\Delta(\lambda)}}{(r + \lambda)^2}$$

which is negative as $xe^{-x} - 1 + e^{-x}$ is negative for any value of $x > 0$.

- When $\lambda < \lambda_p$, we have, using the condition (4) which can be rewritten as follows:

$$\begin{aligned} ye^{-\lambda\Delta} - w &= (r + \lambda) \max [J_p(\lambda), 0] e^{-\lambda\Delta} \\ \frac{e^{-r\Delta(\lambda)} [ye^{-\lambda\Delta} - w]}{r + \lambda} &= \max [J_p(\lambda), 0] e^{-(r+\lambda)\Delta} \end{aligned}$$

Reinserting in (A3) yields

$$J_t(\lambda) = \frac{y}{r + \lambda} - w \left(\frac{1 - e^{-r\Delta(\lambda)}}{r} + \frac{e^{-r\Delta(\lambda)}}{r + \lambda} \right) - c$$

Differentiating yields

$$J'_t(\lambda) = -\frac{y - we^{-r\Delta(\lambda)}}{(r + \lambda)^2} < 0$$

Thus $J_t(\lambda)$ is decreasing in λ everywhere.

A.2 Optimal temporary contract duration in the French system

The optimal duration of temporary contract maximizes the value of starting temporary jobs defined by equation (10). This leads us to distinguish two cases depending whether the contract duration below which contracts are taxed, $\bar{\Delta}$, is either shorter or longer than the optimal contract duration absent taxes defined by

$$\Delta(\lambda, \bar{s}) = \frac{1}{\lambda} \log \left(\frac{w - \bar{s} + \lambda F + (r + \lambda)c}{w - \bar{s}} \right) \quad (\text{A4})$$

Case 1: $\bar{\Delta} \leq \Delta(\lambda, \bar{s})$, the optimal contract duration in the presence of tax maximizes the value of starting temporary jobs defined by the first row of equation (10). This implies that the optimal duration of temporary contracts is given by $\Delta(\lambda, \bar{s})$ as defined by (A4). In this context, the tax has no direct effect on the optimal contract duration for values of the shock arrival rate λ such that the optimal contract duration is longer than the duration $\bar{\Delta}$ below which temporary contracts are taxed. Since $\Delta(\lambda, \bar{s})$ decreases with λ , these values of the shock arrival rate are smaller than the threshold $\bar{\lambda} = \{\lambda | \Delta(\lambda, \bar{s}) = \bar{\Delta}\}$. Moreover, when $\bar{\Delta} \leq \Delta(\lambda, \bar{s})$, the optimal choice of the transformation of temporary contracts into permanent contracts is the same at that defined absent taxes and refunds except that the labor cost is equal to $w - \bar{s}$ instead of w : temporary jobs are transformed into permanent jobs only if $\lambda < \lambda_p^{\text{Fr}} = \{\lambda | J_p(\lambda) = 0\}$. It can be easily checked that $\lambda_p^{\text{Fr}} > \lambda_p$ from the definition (2) of $J_p(\lambda)$.

Case 2: $\bar{\Delta} > \Delta(\lambda)$, two subcases arise. It can be optimal for the firm either to pay the tax and choose the duration, denoted by $\Delta_\tau(\lambda)$, that maximizes $J_t(\lambda, \Delta, \tau) = \int_0^\Delta (ye^{-\lambda z} - w - \tau + \bar{s}) e^{-rz} dz$ or to increase the duration of the contract up to $\bar{\Delta}$ to avoid taxation and get the profit $J_t(\lambda, \bar{\Delta}, 0) = \int_0^{\bar{\Delta}} (ye^{-\lambda z} - w + \bar{s}) e^{-rz} dz$.

Case 2a: If the shock arrival rate is larger than

$$\lambda_\tau = \left\{ \lambda \mid \max_{\Delta} \int_0^\Delta (ye^{-\lambda z} - w - \tau + \bar{s}) e^{-rz} dz = \int_0^{\bar{\Delta}} (ye^{-\lambda z} - w + \bar{s}) e^{-rz} dz \right\}, \quad (\text{A5})$$

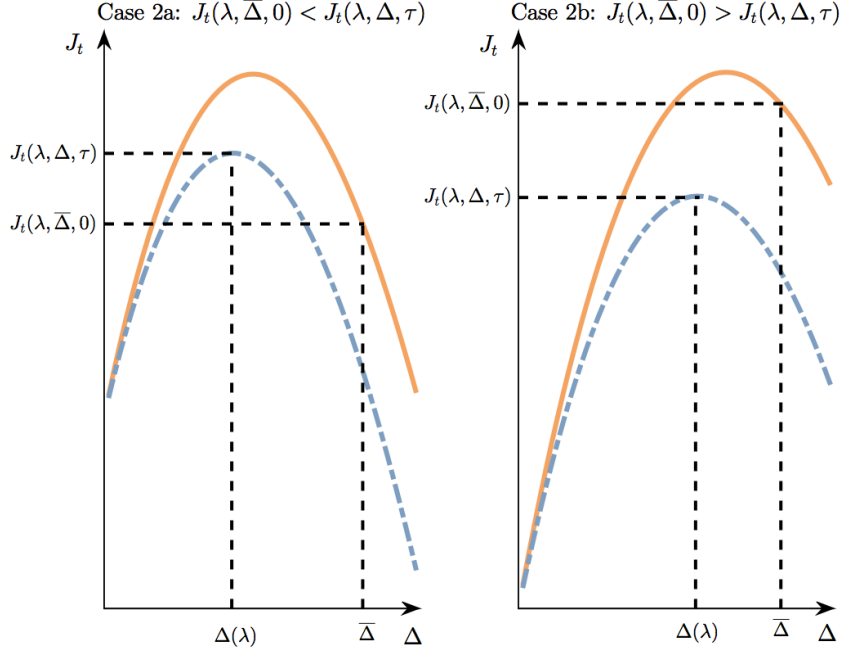


Figure 9: The relation between the value of temporary jobs and the duration of temporary contracts in the French system.

the firm gets higher profits by paying the tax and choosing a duration that maximizes

$$\int_0^{\Delta} \left(y e^{-\lambda x} - w - \tau + \bar{s} \right) e^{-r x} dx.$$

This case is displayed on the left hand side panel of figure 9.

Case 2b: If the shock arrival rate is smaller than λ_{τ} , it is more profitable not to pay the tax and choose the contract duration $\bar{\Delta}$. This case is displayed on the right hand side panel of figure 9.

Finally, the optimal duration of temporary contracts is defined by:

$$\Delta^{\text{Fr}}(\lambda) = \begin{cases} \frac{1}{\lambda} \log \left(\frac{w - \bar{s} + \lambda F + (r + \lambda)c}{w - \bar{s}} \right) & \text{if } \lambda < \lambda_p^{\text{Fr}} \\ \frac{1}{\lambda} \log \left(\frac{y}{w - \bar{s}} \right) & \text{if } \lambda_p^{\text{Fr}} < \lambda \leq \bar{\lambda} \\ \bar{\Delta} & \text{if } \bar{\lambda} < \lambda \leq \lambda_{\tau} \\ \frac{1}{\lambda} \log \left(\frac{y}{w - \bar{s} + \bar{\tau}} \right) & \text{if } \lambda_{\tau} < \lambda \end{cases} \quad (\text{A6})$$

where $\bar{\lambda} = \{\lambda | \Delta(\lambda) = \bar{\Delta}\}$ and λ_{τ} is defined by equation (A5). $\Delta^{\text{Fr}}(\lambda)$ is displayed on figure 5.

A.3 Equilibrium unemployment

This appendix presents the computation of the equilibrium unemployment rate absent taxation on temporary contracts. Let us denote by $l(\lambda)$ the mass of permanent jobs with shock arrival rate λ , by $s_t(\lambda)$ the mass of temporary contracts with shock arrival rate λ that can be converted into permanent contracts and by $s_n(\lambda)$ the mass of temporary contracts with shock arrival rate λ that cannot be converted into permanent contracts.

For all $\lambda \in [\lambda_{\min}, \lambda_s]$, only permanent contracts are created. There are $\alpha ug(\lambda)$ entries into permanent contracts and $\lambda l(\lambda)$ exits. In steady state, we have:

$$\alpha ug(\lambda) = \lambda l(\lambda) \quad (\text{A7})$$

For all $\lambda \in (\lambda_s, \lambda_p]$, only temporary contracts are created and they are transformed into permanent contracts if they are still productive at the end of their spell. The steady state flow equilibrium can be written

$$\alpha ug(\lambda) = \frac{s_t(\lambda)}{\Delta(\lambda)} \quad (\text{A8})$$

$$\frac{s_t(\lambda)}{\Delta(\lambda)} e^{-\lambda \Delta(\lambda)} = \lambda l(\lambda) \quad (\text{A9})$$

For all $\lambda \in (\lambda_p, \lambda_t]$, only temporary contracts are created and they are never transformed into permanent contracts. The steady state flow equilibrium can be written

$$\alpha ug(\lambda) = \frac{s_n(\lambda)}{\Delta(\lambda)} \quad (\text{A10})$$

By definition, the unemployment rate is defined by the following equation:

$$u = 1 - \int_{\lambda_{\min}}^{\lambda_p} l(\lambda) d\lambda - \int_{\lambda_s}^{\lambda_p} s_t(\lambda) d\lambda - \int_{\lambda_p}^{\lambda_t} s_n(\lambda) d\lambda \quad (\text{A11})$$

Using equations (A7) to (A11) we get equation (12).

A.4 GMM estimation

The distribution of contract durations is estimated from the sample (d_1, \dots, d_N) where d_i stands for the duration of contract $i = 1, \dots, N$. It is assumed that the data come from a statistical model defined up to an unknown vector θ of M parameters. Let us denote by $p(s|\theta)$ the share of contracts of spell equal to s days predicted by the model conditional on the vector of parameters θ . Let us define, for

each value s , the indicator function $f(d_i, s) = \mathbf{1}(d_i = s)$ which takes value one for all contract i of spell equal to s and to zero for the others. For each spell s the moment condition is

$$\mathbb{E}[f(d_i, s) - p(s|\theta)] = 0$$

The sample counterpart of $\mathbb{E}[f(d_i, s)]$ is $\frac{1}{N} \sum_{i=1}^N [f(d_i, s)] \equiv p(s)$. Let us assume that there are $S > M$ moment conditions, corresponding to S values of s . The GMM estimator for these S moment conditions is obtained in 2 steps:

1. Let us first define the estimator

$$\hat{\theta} = \arg \min_{\theta} \sum_{s=1}^S [p(s) - p(s|\theta)]^2.$$

This estimator allows us to compute the variance covariance matrix

$$\hat{\Omega} = \frac{1}{N} \sum_{i=1}^N \begin{bmatrix} f(d_i, 1) - p(1|\hat{\theta}) \\ \dots \\ f(d_i, S) - p(S|\hat{\theta}) \end{bmatrix} \begin{bmatrix} f(d_i, 1) - p(1|\hat{\theta}) \\ \dots \\ f(d_i, S) - p(S|\hat{\theta}) \end{bmatrix}'$$

The terms of the diagonal are

$$\frac{1}{N} \sum_{i=1}^N [f(d_i, s) - p(s|\hat{\theta})]^2 = p(s) - 2p(s)p(s|\hat{\theta}) + [p(s|\hat{\theta})]^2$$

and the terms outside the diagonal are, for all $s \neq s'$

$$\frac{1}{N} \sum_{i=1}^N f(d_i, s)f(d_i, s') - p(s)p(s'|\hat{\theta}) - p(s|\hat{\theta})p(s') + p(s|\hat{\theta})p(s'|\hat{\theta})$$

We have $f(d_i, s') = 0$ if $f(d_i, s) = 1$ for all $s \neq s'$ since $f(d_i, s) = 1$ means that the duration of contract i is equal to s , and the same spell cannot be equal to s and to s' . Therefore, we have

$$\frac{1}{N} \sum_{i=1}^N f(d_i, s)f(d_i, s') = 0$$

which implies that the terms outside the diagonal are

$$m(s, s') \equiv p(s|\hat{\theta})p(s'|\hat{\theta}) - p(s)p(s'|\hat{\theta}) - p(s|\hat{\theta})p(s')$$

Since $m(s, s') = m(s', s)$, the variance covariance matrix is

$$\hat{\Omega} = \begin{bmatrix} p(1) - 2p(1)p(1|\hat{\theta}) + [p(1|\hat{\theta})]^2 & m(1, 2) & \dots & m(1, S) \\ m(1, 2) & \dots & \dots & \dots \\ \dots & \dots & \dots & \dots \\ m(1, S) & \dots & \dots & p(S) - 2p(S)p(S|\hat{\theta}) + [p(S|\hat{\theta})]^2 \end{bmatrix}$$

2. The GMM estimator is

$$\hat{\theta} = \arg \min_{\theta} \begin{bmatrix} p(1) - p(1|\theta) \\ \dots \\ p(S) - p(S|\theta) \end{bmatrix}' \hat{\Omega}^{-1} \begin{bmatrix} p(1) - p(1|\theta) \\ \dots \\ p(S) - p(S|\theta) \end{bmatrix}$$

The formula for the variance of the GMM estimator is

$$V(\hat{\theta}) = \frac{1}{N} [G' F^{-1} G]^{-1}$$

where G is the matrix of partial derivatives

$$G = \begin{bmatrix} \frac{\partial p(1|\hat{\theta})}{\partial \theta_1} & \frac{\partial p(1|\hat{\theta})}{\partial \theta_2} & \dots & \frac{\partial p(1|\hat{\theta})}{\partial \theta_M} \\ \frac{\partial p(2|\hat{\theta})}{\partial \theta_1} & \frac{\partial p(2|\hat{\theta})}{\partial \theta_2} & \dots & \dots \\ \dots & \dots & \dots & \dots \\ \frac{\partial p(S|\hat{\theta})}{\partial \theta_1} & \frac{\partial p(S|\hat{\theta})}{\partial \theta_2} & \dots & \frac{\partial p(S|\hat{\theta})}{\partial \theta_M} \end{bmatrix}$$

and F is the *sample* covariance matrix of the moments

$$F = \frac{1}{N} \begin{bmatrix} \sum_{i=1}^N [f(d_i, 1) - p(1)]^2 & \dots & \sum_{i=1}^N [f(d_i, 1) - p(1)] [f(d_i, S) - p(S)] \\ \sum_{i=1}^N [f(d_i, 1) - p(1)] [f(d_i, 2) - p(2)] & \dots & \dots \\ \dots & \dots & \dots \\ \sum_{i=1}^N [f(d_i, 1) - p(1)] [f(d_i, S) - p(S)] & \dots & \sum_{i=1}^N [f(d_i, S) - p(S)]^2 \end{bmatrix}$$

We have

$$\sum_{i=1}^N [f(d_i, s) - p(s)]^2 = Np(s) [1 - p(s)]$$

and, for all $s \neq s'$:

$$\begin{aligned} & \sum_{i=1}^N [f(d_i, s) - p(s)] [f(d_i, s') - p(s')] \\ = & \sum_{i=1}^N f(d_i, s) [f(d_i, s') - p(s')] - \sum_{i=1}^N p(s) \underbrace{[f(d_i, s') - p(s')]_{=0}} \\ = & \sum_{i=1}^N \underbrace{f(d_i, s) f(d_i, s')}_{=0} - p(s') \sum_{i=1}^N f(d_i, s) \\ = & -Np(s')p(s) \end{aligned}$$

therefore, we have

$$F = \begin{bmatrix} p(1) [1 - p(1)] & -p(1)p(2) & \dots & -p(1)p(S) \\ -p(1)p(2) & p(2) [1 - p(2)] & \dots & \dots \\ \dots & \dots & \dots & \dots \\ -p(1)p(S) & \dots & \dots & p(S) [1 - p(S)] \end{bmatrix}$$

The model is overidentified as there are 45 moments and 5 parameters. In order to evaluate the overall match between the model and the data, we use a simple over-identification test *à la* Hansen (1982). Let N be the size of the sample. The statistic

$$N \left[\mathbf{p} - \mathbf{p}(\hat{\boldsymbol{\theta}}) \right] \hat{\Omega}^{-1} \left[\mathbf{p} - \mathbf{p}(\hat{\boldsymbol{\theta}}) \right]'$$

where $\mathbf{p} = (p(1), \dots, p(S))$, $\mathbf{p}(\boldsymbol{\theta}) = (p(1|\boldsymbol{\theta}), \dots, p(S|\boldsymbol{\theta}))$, tests the global adequacy of the model and is asymptotically $\chi^2(T - R)$ distributed. The model is not rejected if the statistic is lower than the critical value of $\chi^2(T - R)$ where T denotes the number of moments and R the number of parameters. Results are reported in Table 1.

Table 1: GMM estimation.

Estimated parameters		
Parameter	Notation	Value
Weibull scale	μ	1.0881 (0.0155)
Weibull shape	σ	0.1841 (0.0033)
Firing costs	F	64.0750 (0.0011)
Cost to write a contract	c	0.0008 (0.0001)
Productivity	y	1.3951 (0.0364)
Baseline parameters		
Discount rate	r	0.000135
Wage	w	1
Job arrival rate	α	0.0228
Hansen statistic		0.1029 $\chi^2 (45 - 5)$

Note: Standard errors in parentheses.

Table 2: Impact of the tax in the French system.

$\tau(\%)$	unemp. rate (%)	Welfare (%)	Mean duration of Temp. jobs			
			< 10 days	< 20 days	< 30 days	all
0	13.5000	0.0000	1.2762	2.1936	2.9926	11.9284
2.5	13.5008	-0.0041	1.1816	2.0310	2.9528	11.8970
5.0	13.5017	-0.0086	1.0892	1.9085	2.9097	11.8621
7.5	13.5027	-0.0133	0.9990	1.8613	2.8642	11.8249
10	13.5037	-0.0184	0.9109	1.8118	2.8164	11.7856

This table presents the impact of the tax on temporary jobs on the unemployment rate, on the welfare of unemployed workers and on the mean duration of temporary jobs. The measure of welfare change is the percentage change in the unemployment benefits replacement ratio equivalent to the change in welfare of unemployed workers induced by the tax. For instance, the tax equal to 10 percent of the labor cost induces a decrease in welfare equivalent to that induced by a drop of 0.0184 percent of the unemployment benefits replacement ratio.

Table 3: Impact of the tax in the Italian system.

τ (%)	unemp. rate (%)	Welfare (%)	Mean duration of Temp. jobs			
			< 10 days	< 20 days	< 30 days	all
0	13.5000	0.0000	1.2762	2.1936	2.9926	11.9284
2.5	13.5095	-0.0454	1.1819	2.0316	2.7716	8.4284
5.0	13.5177	-0.0852	1.0896	1.8731	2.5555	6.8548
7.5	13.5254	-0.1223	0.9992	1.7180	2.3440	5.7420
10	13.5327	-0.1578	0.9102	1.5657	2.1366	4.8696

This table presents the impact of the tax on temporary jobs on the unemployment rate, on the welfare of unemployed workers and on the mean duration of temporary jobs. The measure of welfare change is the percentage change in the unemployment benefits replacement ratio equivalent to the change in welfare of unemployed workers induced by the tax. For instance, the tax equal to 10 percent of the labor cost induces a decrease in welfare equivalent to that induced by a drop of 0.1578 percent of the unemployment benefits replacement ratio.

Table 4: Estimates of the parameters of the matching function.

	(1)	(2)
	OLS	IV
Dep. var.		Labor market tightness (log)
		First stage
Entries		1.51*** (.20)
Entries (-1)		-1.32*** (.19)
Entries (-2)		-0.68** (.28)
Dep. var		Job finding rate (log)
		Second stage
Labor market tightness (log)	.38*** (.07)	.50*** (.16)
Date FE	Yes	Yes
R^2	0,33	
Nb. Observations	879	879

Source : Pôle emploi and EMMO-DMMO. Note: Estimation of the parameter of the job matching function equation (A15) on 348 employment pools from 2005 to 2010. Labor market tightness (log) stands for the first difference in the log of the labor market tightness. Job finding rate (log) stands for the first difference in the log of the job finding rate.

(1) Standard OLS ; (2) IV regression. As instruments we include commuting zone fixed effects and we use Bartik type instrument described in the text. ‘Entries’ stands for the national growth rate of the number of entries into employment across industries using commuting zone industry entries shares as weights. ‘Entries (-1)’ and ‘Entries (-2)’ are the one year and two year lagged values of ‘Entries’ respectively. Robust standard errors in parentheses. * significant at 10 percent, ** significant at 5 percent, *** significant at 1 percent.

Table 5: Impact of the tax in the French system when the labor market tightness is endogenous.

τ (%)	unemp. rate (%)	Welfare (%)	unemp. rate (%)	Welfare (%)	unemp. rate (%)	Welfare (%)
	$\eta = 0.4, \varepsilon_w^l = 0.39$		$\eta = 0.5, \varepsilon_w^l = 1$		$\eta = 0.6, \varepsilon_w^l = 2.73$	
0	13.5000	0.0000	13.5000	0.0000	13.5000	0.0000
2.5	13.5118	-0.0161	13.5188	-0.0349	13.5408	-0.0831
5.0	13.5188	-0.0309	13.5351	-0.0659	13.5815	-0.1558
7.5	13.5252	-0.0450	13.5500	-0.0945	13.6184	-0.2220
10	13.5313	-0.0586	13.5637	-0.1213	13.6522	-0.2833

This table presents the impact of the tax on temporary jobs on the unemployment rate and on the welfare of unemployed workers when the labor market tightness is endogenous for different values of the elasticity of the matching function. The measure of welfare change is the percentage change in the unemployment benefits replacement ratio equivalent to the change in welfare of unemployed workers induced by the tax. For instance, when $\eta = 0.5$, the tax of 10 percent induces a drop of welfare equivalent to that induced by a drop of 0.1213 percent of the unemployment benefits replacement ratio.

Table 6: Impact of the tax in the Italian system when the labor market tightness is endogenous.

τ (%)	unemp. rate (%)	Welfare (%)	unemp. rate (%)	Welfare (%)	unemp. rate (%)	Welfare (%)
	$\eta = 0.4, \varepsilon_w^l = 0.39$		$\eta = 0.5, \varepsilon_w^l = 1$		$\eta = 0.6, \varepsilon_w^l = 2.73$	
0	13.5000	0.0000	13.5000	0.0000	13.5000	0.0000
2.5	13.5421	-0.1008	13.5793	-0.1713	13.6803	-0.3552
5.0	13.5774	-0.1924	13.6535	-0.3298	13.8577	-0.6904
7.5	13.6110	-0.2787	13.7250	-0.4812	14.0307	-1.0157
10	13.6432	-0.3613	13.7942	-0.6273	14.2004	-1.3335

This table presents the impact of the tax on temporary jobs on the unemployment rate and on the welfare of unemployed workers when the labor market tightness is endogenous for different values of the elasticity of the matching function. The measure of welfare change is the percentage change in the unemployment benefits replacement ratio equivalent to the change in welfare of unemployed workers induced by the tax. For instance, when $\eta = 0.5$, the tax of 10 percent induces a drop of welfare equivalent to that induced by a drop of 0.6273 percent of the unemployment benefits replacement ratio.